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Estimation of conditional risk measures for elliptic distributions.

Véronique Maume-Deschamps Joint work with Didier Rullière and Antoine Usseglio-Carleve.

Lisbon, October 17th, 2018



Conclusion

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**Consider four assets**: iShares Core U.S. Aggregate Bond ETF, PowerShares DB Commodity Index Tracking Fund, WisdomTree Europe SmallCap Dividend Fund and SPDR Dow Jones Industrial Average ETF.



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### A financial example



How to use the knowledge of the 4 variables in order to estimate risk measures for WisdomTree Japan Hedged Equity Fund.



### Consider $(Z_t)_{t \in T}$ a random field observed at *n* points $t_1, ..., t_n$ .



#### Spatial Process

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### A general temporal problematic

### Consider $(Z_t)_{t \in T}$ a random field observed at *n* points $t_1, ..., t_n$ .



#### **Spatial Process**

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### A general temporal problematic

Consider  $(Z_t)_{t \in T}$  a random field observed at *n* points  $t_1, ..., t_n$ . Using kriging ([Krige, 1951], [Matheron, 1963]) lead to estimations of the conditional mean, reliable for Gaussian processes.



How to estimate conditionnal risk measures (quantiles, expectiles, TVaR), for non necessarily gaussian random fields ?



Rainfall amounts measured at some locations  $X \in \mathbb{R}$ , one unknown value  $Y \in \mathbb{R}$ .



Source: Geographic Information Technology Training Alliance. How to estimate risk measures related to Y knowing X?

Introduction	Background	<b>Regression methods</b>	High level risk measures	A real data example	Conclusion
Our pu	rpose				

- Focus on elliptic distributions,
- Consider quantile / expectile regression
- Propose an alternative estimation, for high level risk measures.

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Elliptic distribu	tions				
Definit	ion				

#### Definition (Representation Theorem [Cambanis et al., 1981])

A  $\mathbb{R}^d$  random vector X is elliptical with parameters  $\mu$  and  $\Sigma$ (X ~  $\mathcal{E}_d(\mu, \Sigma)$ ) if it writes:

$$X \stackrel{d}{=} \mu + R\Lambda U^{(d)}$$

where  $\Lambda\Lambda^T = \Sigma$ ,  $U^{(d)}$  is a d-dimensional random vector uniformly distributed on  $S^{d-1}$ , R is a non-negative random variable, independent of  $U^{(d)}$ .

• *R* is called the radius of *X*.

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**Elliptic distributions** 

### Consistent Elliptic distributions

Background

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Definition (Kano, 1994 [Kano, 1994])

 $X \sim \mathcal{E}_d(\mu, \Sigma, R)$  has the consistency property if:

$$R \stackrel{d}{=} \chi_d \epsilon$$

where  $\epsilon$  is a positive random variable unrelated to d and independent of  $\chi_d$ . We shall say that X is a consistent (R, d)-elliptical random vector with parameters  $\mu$  and  $\Sigma$ .

Remark: the above definition rewrites as

$$X \stackrel{d}{=} \mu + \epsilon \mathcal{N}(0, \Sigma)$$

with  $\epsilon$  independent of the normal vector. This means that finally, an elliptical distribution is a normal distribution with random variance  $\epsilon^2 \Sigma$ ). 



 Sub-vectors of elliptical vectors are elliptical, more precisely, Let X = (X<sub>1</sub>, X<sub>2</sub>) be a consistent (R, d)-elliptical random vector with parameters μ and Σ. X<sub>1</sub> and X<sub>2</sub> are d<sub>1</sub> and d<sub>2</sub>-dimensional subvectors of X. Let us write Σ :

$$\Sigma = egin{pmatrix} \Sigma_{11} & \Sigma_{12} \ \Sigma_{21} & \Sigma_{22} \end{pmatrix}$$

Then  $X_1$  and  $X_2$  are respectively  $(R, d_1)$ - and  $(R, d_2)$ -elliptical with parameters  $\mu_1$ ,  $\Sigma_{11}$  and  $\mu_2$ ,  $\Sigma_{22}$ , respectively.

• Conditional distributions of elliptical vectors are also elliptical.



- Sub-vectors of elliptical vectors are elliptical,
- Conditional distributions of elliptical vectors are also elliptical. More precisely,

 $X_2|(X_1 = x_1)$  is still elliptical, with radius  $R^*$  given by:

$$R^* \stackrel{d}{=} R\sqrt{1-\beta} | \left( R\sqrt{\beta} U^{(d)} = C_{11}^{-1}(x_1-\mu_1) \right),$$

where  $C_{11}$  is the Cholesky root of  $\Sigma_{11}$ , and  $\beta \sim Beta(\frac{d_1}{2}, \frac{d_2}{2})$ .

Conditional measures.



- Sub-vectors of elliptical vectors are elliptical,
- Conditional distributions of elliptical vectors are also elliptical.

#### Definition

A random field  $(Z_t)_{t\in T}$ ,  $t\in \mathbb{R}$  is  $\epsilon$ -elliptical if

$$Z_t = \mu + \epsilon X_t$$

where  $(X_t)_{t \in T}$  is a gaussian field.

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Exampl	es				

• Normal distributions:  $\epsilon = 1$ .



Densité gaussienne

- Student distributions: with  $\nu$  degrees of freedom:  $\epsilon \stackrel{d}{=} \sqrt{\frac{\nu}{\chi_d^2}}$ .
- Slash distributions:  $\epsilon \stackrel{d}{=} \mathcal{P}(1, a)$ .
- Laplace distibution:  $\epsilon \stackrel{d}{=} \sqrt{\mathcal{E}(\lambda)}$ .
- Many other.

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Elliptic distribut	tions				
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- Normal distributions:  $\epsilon = 1$ .
- Student distributions: with  $\nu$  degrees of freedom:  $\epsilon \stackrel{d}{=} \sqrt{\frac{\nu}{\chi_d^2}}$ .

Densité Student



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- Slash distributions:  $\epsilon \stackrel{d}{=} \mathcal{P}(1, a)$ .
- Laplace distibution:  $\epsilon \stackrel{d}{=} \sqrt{\mathcal{E}(\lambda)}$ .
- Many other.

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Examp	les				

- Normal distributions:  $\epsilon = 1$ .
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- Laplace distibution:  $\epsilon \stackrel{d}{=} \sqrt{\mathcal{E}(\lambda)}$ .
- Many other.

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- Normal distributions:  $\epsilon = 1$ .
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- Slash distributions:  $\epsilon \stackrel{d}{=} \mathcal{P}(1, a)$ .
- Laplace distibution:  $\epsilon \stackrel{d}{=} \sqrt{\mathcal{E}(\lambda)}$ .

Densité Laplace





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Quantile is widely used as a risk measure (VaR). Recall: for X a random variable (a risk) with distribution function  $F_X$ ,

- $q_{\alpha}(X) = \operatorname{VaR}_{\alpha}(X) = \inf\{t \ / \ F_X(t) \ge \alpha\} = F_X^{-1}(\alpha),$
- RiskMetrics popularized the use of VaR as a risk measure (1994).
- Basel Committee : Internal approach to capital management using VaR (1996),
- formalisation of coherent risk measures [Artzner et al., 1999], VaR is not coherent (because not sub-additive in general).
   VaR is a risk threshold but does not gives information on the risk above the threshold.



TVaR / Expected shortfall: (for continuous distributions) - see e.g. [Artzner et al., 1999] -

$$\begin{aligned} \mathsf{T}\mathsf{VaR}_{\alpha}(X) &= \mathsf{ES}_{\alpha}(X) = \mathbb{E}(X|X > \mathsf{VaR}_{\alpha}(X)) \\ &= \frac{1}{1-\alpha} \int_{\alpha}^{1} q_{u}(X) du. \end{aligned}$$

TVaR is a coherent risk measure.

The notion of ellicitability has taken some importance these last years [Bellini and Di Bernardino, 2017] in risk management.



A statistic T is ellicitable ([Gneiting, 2011, Ziegel, 2016]) if it can be written as a minimizer of a scoring function denoted  $s : \mathbb{R} \times \mathbb{R} \to \mathbb{R}$ :

$$T(X) = \operatorname*{argmin}_{x \in \mathbb{R}} \mathbb{E}[s(X, x)].$$

Interest:

- Ability to compare different statistical methods
- Makes backtesting procedures easier
- Allow the use of stochastic approximation tools.

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VaR is ellicitable: take

$$s(x,X) = [\alpha(X-x)_+ + (1-\alpha)(X-x)_-].$$

TVaR is not ellicitable.

Define the expectile as ([Newey and Powell, 1987]):

$$e_{\alpha}(X) = \operatorname*{argmin}_{x \in \mathbb{R}} \mathbb{E} \left[ lpha(X-x)_{+}^{2} + (1-lpha)(X-x)_{-}^{2} 
ight].$$

Expectiles are:

- Ellicitable (of course),
- Coherent,
- the only risk measure which is both coherent and ellicitable ([Bellini and Bignozzi, 2015]).

Expectiles satisfy

$$\alpha \mathbb{E}\left[(X - e_{\alpha}(X))_{+}\right] = (1 - \alpha) \mathbb{E}\left[(X - e_{\alpha}(X))_{-}\right]$$

and thus take into account the risk over the threshold. 24/57



We are interested in conditional risk measures for elliptical distributions. Focus on quantiles (results also for TVaR and expectiles).

Let  $X = (X_1, X_2)$  be a (R, N + 1)-elliptical random vector with parameters  $\mu$  and  $\Sigma$ .  $X_1 \in \mathbb{R}^N$  and  $X_2 \in \mathbb{R}$ . Calculate / estimate

 $q_{\alpha}(X_2|X_1=x_1), \ e_{\alpha}(X_2|X_1=x_1).$ 

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Condit	ional risk	( measures			

#### Proposition

Let  $X = (X_1, X_2)$  a (R, N + 1)-elliptical random vector with parameters  $\mu$  and  $\Sigma$ . Write

$$\Sigma = egin{pmatrix} \Sigma_{11} & \Sigma_{12} \ \Sigma_{21} & \Sigma_{22} \end{pmatrix}$$

Then

$$q_{\alpha}(X_2|X_1=x_1) = \mu_{2|1} + \sqrt{\Sigma_{2|1}} \Phi_{R^*}^{-1}(\alpha)$$

with  $\begin{cases} \mu_{2|1} = \mu_2 + \Sigma_{21} \Sigma_{11}^{-1} (x_1 - \mu_1) \\ \Sigma_{2|1} = \Sigma_{22} - \Sigma_{21} \Sigma_{11}^{-1} \Sigma_{12} \end{cases}$  and for a random variable Y,  $\Phi_Y$  is the distribution function of  $\frac{Y - \mathbb{E}(Y)}{\sigma_Y}$ .

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$$q_{\alpha}(X_{2}|X_{1}=x_{1})=\mu_{2|1}+\sqrt{\Sigma_{2|1}}\Phi_{R^{*}}^{-1}(\alpha)$$

Problem: the distribution of  $R^*$  is hardly accessible.

- Explicit formulas for Gaussian and Student distributions,
- First idea: quantiles regression,
- Extreme estimations (ie  $\alpha \rightarrow 1$ ).

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Quanti	le Regre	ssion			

Approximate the conditional quantile by: ([Koenker and Bassett, 1978])

$$\hat{q}_{lpha}(X_2|X_1=x_1)={eta^*}^{ op}x_1+{eta_0^*}$$

where  $\beta^*$  and  $\beta^*_0$  are the solutions of the following minimization problem:

$$(\beta^*, \beta_0^*) = \arg\min_{\beta \in \mathbb{R}^N, \beta_0 \in \mathbb{R}} \mathbb{E}[s_\alpha (X_2 - \beta^T X_1 - \beta_0)]$$

with the scoring function  $s_{\alpha}$ :

$$s_{\alpha}(x) = (\alpha - 1)x \mathbf{1}_{\{x < 0\}} + \alpha x \mathbf{1}_{\{x > 0\}}$$

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Quantile regression

### Quantile Regression for elliptic distributions

#### Theorem

Let  $X = (X_1, X_2)$  be an elliptical distribution, the optimal  $\beta^*$  is given by :

$$\beta^* = \Sigma_{11}^{-1} \Sigma_{12}$$

The Quantile Regression Predictor with level  $\alpha \in [0,1]$  is given by:

$$\hat{q}_{lpha}(X_2|X_1=x_1)=\mu_{2|1}+\sqrt{\Sigma_{2|1}}\Phi_R^{-1}(lpha)$$

It satisfies

$$\hat{q}_{lpha}(X_2|X_1) \sim \mathcal{E}_1\left(\mu_2 + \sigma_{2|1}\Phi_R^{-1}(lpha), \Sigma_{21}\Sigma_{11}^{-1}\Sigma_{12}, R
ight)$$

Recall the theoretical conditional quantile:

$$q_{\alpha}(X_{2}|X_{1}=x_{1}) = \mu_{2|1} + \sqrt{\sum_{2|1}} \Phi_{R^{*}(\alpha)}^{-1}(\alpha)$$

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### How good is the quantile regression?

#### Gaussian case

$$\begin{cases} q_{\alpha}(X_{2}|X_{1} = x_{1}) = & \mu_{2|1} + \sigma_{2|1}\Phi^{-1}(\alpha) \\ \hat{q}_{\alpha}(X_{2}|X_{1} = x_{1}) = & \mu_{2|1} + \sigma_{2|1}\Phi^{-1}(\alpha) \end{cases}$$

The Quantile Regression Predictor is exactly the conditional quantile.

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Quantile regression

### How good is the quantile regression?

Student case

$$\begin{cases} q_{\alpha}(X_{2}|X_{1}=x_{1}) = & \mu_{2|1} + \sigma_{2|1}\sqrt{\frac{\nu}{\nu+N}}\sqrt{1 + \frac{1}{\nu}d_{1}}\Phi_{\nu+N}^{-1}(\alpha) \\ \hat{q}_{\alpha}(X_{2}|X_{1}=x_{1}) = & \mu_{2|1} + \sigma_{2|1}\Phi_{\nu}^{-1}(\alpha) \end{cases}$$

The error may be huge, especially if the Mahalanobis distance  $d_1 = (x_1 - \mu_1)^T \Sigma_{11}^{-1} (x_1 - \mu_1)$  is high. Below N = 5.

Student Quantile Regression



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Recall that the  $\alpha$ -expectile of X is defined as:

$$\underset{x \in \mathbb{R}}{\arg\min} \mathbb{E}\left[\left(1-\alpha\right)\left(x-X\right)_{+}^{2}+\alpha\left(X-x\right)_{+}^{2}\right].$$

In the elliptical case, the expectile of level  $\alpha$  is the solution of:

$$\Psi_R(x) = \Phi_R(x) + \frac{1}{x} \int_x^{+\infty} y c_1 g_1(y^2) dy = \frac{\alpha}{2\alpha - 1}.$$

Theoretical conditional expectiles :

$$e_{\alpha}(X_2|X_1=x_1)=\mu_{2|1}+\sigma_{2|1}\Psi_{R^*}^{-1}\left(\frac{lpha}{2lpha-1}
ight).$$

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Approximate the conditional expectile by: ([Newey and Powell, 1987])

$$\hat{e}_{\alpha}(X_2|X_1=x_1) = \beta^{*T}x_1 + \beta_0^*$$

where  $\beta^*$  and  $\beta^*_{\rm 0}$  are the solutions of the minimization problem :

$$(\beta^*, \beta_0^*) = \operatorname*{arg\,min}_{\beta \in \mathbb{R}^N, \beta_0 \in \mathbb{R}} \mathbb{E}[s_\alpha (X_2 - \beta^T X_1 - \beta_0)]$$

with the loss function  $s_{\alpha}$ :

$$s_{\alpha}(x) = (1-\alpha)x^2 \mathbf{1}_{\{x<0\}} + \alpha x^2 \mathbf{1}_{\{x>0\}}$$

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Expectile regression

### Expectile Regression for elliptical distributions

#### Theorem

Let  $X = (X_1, X_2)$  be an elliptical distribution, the optimal  $\beta^*$  is given by :

$$\begin{cases} \beta^* = \Sigma_{11}^{-1} \Sigma_{12} \\ \beta_0^* = \mu_2 - \Sigma_{21} \Sigma_{11}^{-1} \mu_1 + \sigma_{2|1} \Psi_R^{-1} \left( \frac{\alpha}{2\alpha - 1} \right) \end{cases}$$

The Expectile Regression Predictor  $\alpha \in [0,1]$  is:

$$\hat{e}_{\alpha}(X_2|X_1=x_1) = \mu_{2|1} + \sigma_{2|1} \Psi_R^{-1}\left(\frac{\alpha}{2\alpha-1}\right).$$

Furthermore,

$$\hat{e}_{\alpha}(X_2|X_1) \sim \mathcal{E}_1\left(\mu_2 + \sigma_{2|1}\Psi_R^{-1}\left(\frac{\alpha}{2\alpha - 1}\right), \Sigma_{21}\Sigma_{11}^{-1}\Sigma_{12}, R\right).$$



### How good is the expectile regression?

Gaussian case The Quantile Regression Predictor is exactly the conditional quantile.

Student case Semi-explicit formula,  $\Psi_R^{-1}$  computed using MM algorithms. The error may be huge. Below N = 5.





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Extreme approximations

In case  $\alpha \sim 1$  or  $\alpha \sim 0$ , alternative methods have to be proposed. In the case of quantile, we found an equivalent of  $\Phi_{R^*}^{-1}(\alpha)$ . In the case of expectile, we found an equivalent of  $\Psi_{R^*}^{-1}(\alpha)$ . Introduction

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### Some asymptotic relationships

#### Theorem

Under some technical assumptions, their exist  $0 < \ell < +\infty$  and  $\eta \in \mathbb{R}$  such that :

$$\left[\Phi_{R}^{-1}\left(1-\frac{1}{\frac{\ell}{1-\alpha}+2(1-\ell)}\right)\right]^{\frac{1}{\eta}} \underset{\alpha \to 1}{\sim} \Phi_{R^{*}}^{-1}(\alpha)$$

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#### Property

The Gaussian, Student and Slash distributions satisfy the previous assumptions, with coefficients  $\eta$  and  $\ell$  given in the table below.

Distribution	$\eta$	l
Gaussian	1	1
Student, $ u > 0$	$rac{N}{ u}+1$	$rac{\Gamma\left(rac{ u+N+1}{2} ight)\Gamma\left(rac{ u}{2} ight)}{\Gamma\left(rac{ u+N}{2} ight)\Gamma\left(rac{ u+1}{2} ight)}\left(1+rac{q_1}{ u} ight)^{rac{N+ u}{2}}rac{ u}{ u+N}$
Slash, <i>a</i> > 0	$\frac{N}{a} + 1$	$\frac{\Gamma\left(\frac{N+1+a}{2}\right)q_1^{\frac{N+a}{2}}}{\Gamma\left(\frac{N+a}{2}\right)(N+a)\chi^2_{N+a}(q_1)2^{\frac{a}{2}-1}\Gamma\left(\frac{1+a}{2}\right)}$

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### Extremal correction in the Student case

Student Extremal Predictor



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Under some technical assumptions, their exist 0  $<\ell<+\infty$  and  $\gamma\in\mathbb{R}$  such that:

$$\left\{ egin{array}{ll} \hat{\hat{e}}_{lpha\uparrow}(X_2|X_1=x_1) & \sim & e_lpha(X_2|X_1=x_1) \ \hat{\hat{e}}_{lpha\downarrow}(X_2|X_1=x_1) & \sim & e_lpha(X_2|X_1=x_1) \ & & lpha
ightarrow e_lpha(X_2|X_1=x_1) \end{array} 
ight.$$

with

$$\begin{cases} \hat{\hat{e}}_{\alpha\uparrow}(X_2|X_1=x_1) &= \mu_{2|1} + \sigma_{2|1} \left[ \Psi_R^{-1} \left( 1 - \frac{\alpha - 1}{(2\alpha - 1)\ell} \right) \right]^{\frac{1}{\gamma}} \\ \hat{\hat{e}}_{\alpha\downarrow}(X_2|X_1=x_1) &= \mu_{2|1} - \sigma_{2|1} \left[ \Psi_R^{-1} \left( 1 - \frac{\alpha}{(2\alpha - 1)\ell} \right) \right]^{\frac{1}{\gamma}} \end{cases}$$

Hypothesis satisfied for Gaussian, Student, Slash distributions.

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#### Extremal correction in the Student case

Student Extremal Predictor



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Under additional assumptions (heavy tail + order two condition, estimations of the parameters  $\ell$ ,  $\eta$ ,  $\gamma$  + asymptotic normality of the estimators ([Usseglio-Carleve, 2017]).

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## Financial example

Background



[Usseglio-Carleve, 2017]. These four values are the first available every day  $\Rightarrow$  anticipate the behaviour of the return of WisdomTree Japan Hedged Equity Fund  $X_2$ . 

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The sample size is 2520. The first 2519 days (from January 3, 2007 to December 5, 2016) = learning sample, and we focus on the 2520th day:  $x_1 = (-0.0185\%, -0.4464\%, 0.9614\%, 0.1405\%)$ . Estimate quantiles / expectiles / TVaR of  $X_2|X_1 = x_1$ .

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The sample size is 2520. The first 2519 days (from January 3, 2007 to December 5, 2016) = learning sample, and we focus on the 2520th day:  $x_1 = (-0.0185\%, -0.4464\%, 0.9614\%, 0.1405\%)$ . Estimate quantiles / expectiles / TVaR of  $X_2 | X_1 = x_1$ . Data exploration:

- the daily returns can be considered as independent.
- the marginals seem symmetrical.
- the measured tail index is approximately the same for the marginals.

Could be assumed to be elliptical.

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E.g., for  $\alpha$  = 0.999, the estimated VaR is 3.1%, the observed value is 0.7141%.

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### Conclusion / perspectives

- Regression methods are not satisfactory for non gaussian distributions.
- Framework adapted to a large class of risk measures (*L<sup>p</sup>* quantile, Haezendonck-Goovaerts risk measures).
- New technics needed in the high dimension case (*N* large), see Antoine Usseglio-Carleve (2019).
- More details in [Maume-Deschamps et al., 2017a, Maume-Deschamps et al., 2017b, Usseglio-Carleve, 2017].
- Mixed approaches for non central but non extreme risk levels?
- Non symetric distributions?

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Thank you	Introduction	Background	Regression methods	High level risk measures	A real data example	Conclusion
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# Obrigada pela vossa atenção.