

# Parameters estimation of a Threshold CIR process

*Benoît Nieto* (ICJ,ECL Lyon)

travail commun avec *Sara Mazzonetto* (IECL/Pasta, Inria Nancy)

Colloque Jeunes Probabilistes et Statisticiens - 2023

- 1 Introduction
- 2 Cox–Ingersoll–Ross (CIR) model
  - Model
  - Literature
- 3 Threshold Cox-Ingersoll-Ross (TCIR) model
  - Existence of a strong solution and property
  - Drift Estimations from continuous observations
  - Drift Estimations from high frequency and long time observations
- 4 Conclusion and Opening

- 1 Introduction
- 2 Cox–Ingersoll–Ross (CIR) model
  - Model
  - Literature
- 3 Threshold Cox-Ingersoll-Ross (TCIR) model
  - Existence of a strong solution and property
  - Drift Estimations from continuous observations
  - Drift Estimations from high frequency and long time observations
- 4 Conclusion and Opening

# Introduction

The Cox–Ingersoll–Ross (CIR) model is a stochastic process that exhibits mean-reverting behaviour.

- In finance, the CIR is used to describes the evolution of interest rates.
- In biology, the CIR can be used to model population dynamics [Bansaye & Méléard, 2015].

Threhsold Diffusion:

- Threshold autoregressive (TAR) models in discrete time were introduced in the early 1980s.
- Many applications in Finance, Physics and meteorology.

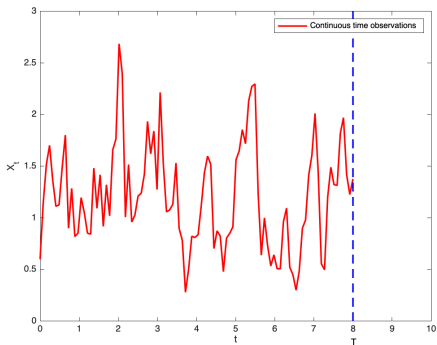
The Threshold CIR follows the CIR dynamics when above or below a fixed level, yet at this level (threshold) its coefficients can be discontinuous.

**Goal:** Estimate the parameters of a TCIR process by using the observations of a single trajectory.

- MLE drift estimation.

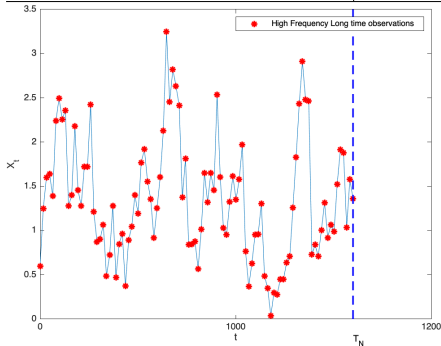
→ Study of the asymptotic behavior of the estimators (Consistency, Asymptotic Normality).

- Continuous time observations.
- High frequency and long time observations.



## Continuous time observations:

We observe the process on the interval  $[0, T]$  with  $T \in (0, \infty)$ .



## High Frequency in Long time observations:

We observe the process on the discrete time grid

$$0 < t_0 < \dots < t_N = T, \text{ for } N \in \mathbb{N} \text{ and } T \in (0, \infty).$$

- 1 Introduction
- 2 Cox–Ingersoll–Ross (CIR) model
  - Model
  - Literature
- 3 Threshold Cox-Ingersoll-Ross (TCIR) model
  - Existence of a strong solution and property
  - Drift Estimations from continuous observations
  - Drift Estimations from high frequency and long time observations
- 4 Conclusion and Opening

We work on the probability space  $(\Omega, (\mathcal{F}_t)_{t \geq 0}, \mathbb{P})$ .

## Cox–Ingersoll–Ross (CIR) process

Let  $(X_t)_{t \geq 0}$ , the process solution of

$$\begin{cases} dX_t = (a - bX_t)dt + \sqrt{2\sigma X_t}dB_t, & t > 0, \\ X_0 = x_0, \end{cases} \quad (1)$$

with  $(a, b, \sigma) \in \mathbb{R}_+ \times \mathbb{R} \times \mathbb{R}_+^*$ ,  $(B_t)_{t \geq 0}$  a standard Brownian motion.

Strong existence of a unique solution to (1) follows from the classical results in 1-D.

## Property

- For all  $t \geq 0$ , the process  $(X_t)_{t \geq 0}$  is positif (Comparison Theorem).
- If  $a \leq \sigma$ ,  $\{0\}$  instantaneously reflecting.
- If  $a > \sigma$ , the state space of the process is  $]0, +\infty[$ .



In [Alaya & Kebaier, 2013], the authors study the Maximum Likelihood Estimator for the parameters  $(a, b)$ .

## Likelihood ration/Girsanov Weight

Suppose that the one dimensional diffusion process  $(X_t^\theta)_{t \geq 0}$  satisfies

$$dX_t^\theta = b(\theta, X_t)dt + \sigma(X_t)dB_t, \quad X_0 = x_0.$$

- Continuous observations  $(X_t^{\theta_0})_{t \in [0, T]}$ .
- $\theta_0 \in \Theta$  the parameter to be estimated.

For any  $\theta \in \Theta$ , the Likelihood ratio is given by

$$G_T^{\theta, \theta_0} := \frac{d\mathbb{P}_{\theta, T}}{d\mathbb{P}_{\theta_0, T}} = \exp \left( \int_0^T \frac{b(\theta, X_s) - b(\theta_0, X_s)}{\sigma^2(X_s)} dX_s - \frac{1}{2} \int_0^T \frac{b^2(\theta, X_s) - b^2(\theta_0, X_s)}{\sigma^2(X_s)} ds \right).$$

For the CIR process the Likelihood ratio, evaluated at time  $T$  is given by

$$G_T(\theta) = G_T(a, b) := G_T^{\theta, \theta_0} = \exp \left( \frac{1}{2\sigma} \int_0^T \frac{a - bX_s}{X_s} dX_s - \frac{1}{4\sigma} \int_0^T \frac{(a - bX_s)^2}{X_s} ds \right).$$

$$(\hat{a}_T, \hat{b}_T) = \underset{(a,b) \in \Theta}{\operatorname{Argmax}} G_T(a, b) \text{ (Explicitly computable).}$$

### Remark

The estimators and the Likelihood ratio is well defined iff  $a > \sigma$  i.e.

$$\mathbb{P}_\theta \left( \int_0^T \frac{ds}{X_s} < \infty \right) = 1.$$

So  $\Theta \in \{(a, b) \text{ s.a. } a > \sigma\}$ .

Study of the asymptotic behavior in the ergodic case ( $b > 0$ ) and non ergodic case ( $b \leq 0$ ).

- in long time from continuous time observations.
- in long time from high frequency observations (for any  $x_0 \in \mathbb{R}^+$ ).

- 1 Introduction
- 2 Cox–Ingersoll–Ross (CIR) model
  - Model
  - Literature
- 3 Threshold Cox-Ingersoll-Ross (TCIR) model
  - Existence of a strong solution and property
  - Drift Estimations from continuous observations
  - Drift Estimations from high frequency and long time observations
- 4 Conclusion and Opening

## Threshold Cox-Ingersoll-Ross (TCIR) model

Let  $(X_t)_{t \geq 0}$ , the process solution of

$$\begin{cases} dX_t = (a_r(X_t) - b_r(X_t)X_t)dt + \sqrt{2\sigma_r(X_t)X_t}dB_t, & t > 0, \\ X_0 = x_0. \end{cases} \quad (2)$$

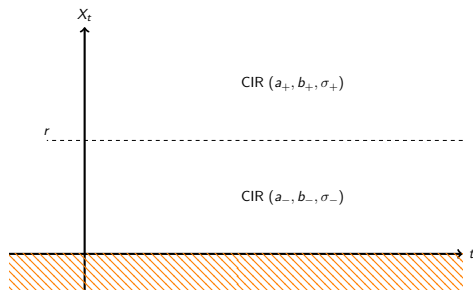
$$a_r(x) = \begin{cases} a_+ \in \mathbb{R} & \text{si } x \geq r, \\ a_- \in \mathbb{R}^+ & \text{si } x < r, \end{cases} \quad b_r(x) = \begin{cases} b_+ \in \mathbb{R} & \text{si } x \geq r, \\ b_- \in \mathbb{R} & \text{si } x < r, \end{cases}$$
$$\sigma_r(x) = \begin{cases} \sigma_+ > 0 & \text{si } x \geq r, \\ \sigma_- > 0 & \text{si } x < r. \end{cases} \quad (3)$$

with  $r > 0$  et  $x_0 > 0$ .

**Issue:** Drift and volatility does not fit into the classical hypotheses of strong existence and unicity of a dimensional EDS.

→ Strong existence of a unique solution to (1) follows from the results of [Mazzonetto & Nieto, in progress].

# Property of the TCIR



Separately on  $(r, \infty)$  and  $(0, r)$ , the process follows the CIR dynamics.

## Property

- For all  $t \geq 0$ , the process  $(X_t)_{t \geq 0}$  is positif (Comparison Theorem [Mazonetto & Nieto, in progress]).
- If  $a_- \leq \sigma_-$ ,  $\{0\}$  instantaneously reflecting.
- If  $a_- > \sigma_-$ , the state space of the process is  $]0, +\infty[$ .

# Regime of the process

We denote the scale function  $S$  and speed measure  $m$ :

$$S'(x) = \exp \left( \int^x \frac{(a_r(y) - b_r(y)y)}{\sigma_r(y)y} dy \right) \text{ and } m(x) = \frac{1}{\sigma_r(x)xS'(x)}.$$

Long time behavior depends on the parameters:

	Null recurrent	Positive Recurrent (Ergodic)	Transience
$\{0\}$ Absorbing point ( $a_- = 0$ )	$[b_+ = 0 \text{ and } 0 \leq a_+ < \sigma_+]$ and $[b_- \in \mathbb{R} \text{ and } a_- \leq 0]$	$\emptyset$	$[b_+ < 0 \text{ or } a_+ \geq \sigma_+]$
$\{0\}$ Instantaneous reflecting point ( $a_- \leq \sigma_-$ )	$[b_+ = 0 \text{ and } 0 \leq a_+ < \sigma_+]$ and $[b_- \in \mathbb{R} \text{ and } 0 < a_- \leq \sigma_-]$ .	$[[b_+ > 0 \text{ and } a_+ \in \mathbb{R}]$ or $[b_+ = 0 \text{ and } a_+ \in \mathbb{R}^*]]$ and $[b_- \in \mathbb{R} \text{ and } 0 < a_- \leq \sigma_-]$ .	$[b_+ < 0 \text{ or } a_+ \geq \sigma_+]$
$\{0\}$ is not reached ( $a_- > \sigma_-$ )	$[b_+ = 0 \text{ and } 0 \leq a_+ < \sigma_+]$ and $[b_- \in \mathbb{R} \text{ and } a_- > \sigma_-]$	$[[b_+ > 0 \text{ and } a_+ \in \mathbb{R}]$ or $[b_+ = 0 \text{ and } a_+ \in \mathbb{R}^*]]$ and $[b_- \in \mathbb{R} \text{ and } a_- > \sigma_-]$ .	$[b_+ < 0 \text{ or } a_+ \geq \sigma_+]$

In the positive recurrent case, the stationary measure  $\mu$ , is given by

$$\mu(dx) = \frac{m(x)}{\int_0^{+\infty} m(y)dy} dx.$$

We denote  $\theta = (a_+, b_+, a_-, b_-)$ , the likelihood ratio evaluated at time  $T$  is given by

$$G_T(\theta) = \exp \left( \int_0^T \frac{a_r(X_s) - b_r(X_s)X_s}{2\sigma_r(X_s)X_s} dX_s - \frac{1}{4} \int_0^T \frac{(a_r(X_s) - b_r(X_s)X_s)^2}{\sigma_r(X_s)X_s} ds \right).$$

## Remark

The estimators and the Likelihood is well defined iff  $a_- > \sigma_-$  i.e.

$$\mathbb{P}_\theta \left( \int_0^T \frac{1_{X_s \leq r} ds}{X_s} < \infty \right) = 1.$$

**For the estimation problem, we suppose  $r$  and  $\sigma_\pm$  to be known.**

# Drift estimation from continuous time observations

Imagine that we observe  $(X_t)_{t \in [0, T]}$  in continuous time. For  $T \in (0, \infty)$  and  $m = -1, 0, 1$ , we define

$$\mathcal{Q}_T^{\pm, m} = \int_0^T X_s^m 1_{\{\pm(X_s - r) \geq 0\}} ds \quad \text{and} \quad \mathcal{M}_T^{\pm, m} = \int_0^T X_s^m 1_{\{\pm(X_s - r) \geq 0\}} dX_s.$$

Let  $\theta_0 = (a_0^{\pm}, b_0^{\pm})$ , the parameter to be estimated and  $\theta \in \Theta \subset \mathbb{R}^4$ . We denote

$$(\alpha_T^{\pm}, \beta_T^{\pm}) = \underset{\theta \in \Theta}{\operatorname{Argmax}} G_T(\theta).$$

## Estimators

For every  $T \in (0, \infty)$  the MLE are given by

$$\alpha_T^{\pm} = \frac{\mathcal{M}_T^{\pm, -1} \mathcal{Q}_T^{\pm, 1} - \mathcal{M}_T^{\pm, 0} \mathcal{Q}_T^{\pm, 0}}{\mathcal{Q}_T^{\pm, 1} \mathcal{Q}_T^{\pm, -1} - (\mathcal{Q}_T^{\pm, 0})^2} \quad \text{and} \quad \beta_T^{\pm} = \frac{\mathcal{M}_T^{\pm, -1} \mathcal{Q}_T^{\pm, 0} - \mathcal{Q}_T^{\pm, -1} \mathcal{M}_T^{\pm, 0}}{\mathcal{Q}_T^{\pm, 1} \mathcal{Q}_T^{\pm, -1} - (\mathcal{Q}_T^{\pm, 0})^2},$$



# Theorem MLE: long-time behavior (positive recurrent cases)

Let  $\pm \in \{+, -\}$ . It holds that

- $\frac{1}{T} Q_T^{\pm, m} \xrightarrow[T \rightarrow \infty]{\text{a.s.}} Q_{\infty}^{\pm, m} \in \mathbb{R}_+^*$  (we have explicit expressions using Ergodic Theorem).
- The estimator is strongly consistent  $(\alpha_T^{\pm} - a_{\pm}, \beta_T^{\pm} - b_{\pm}) \xrightarrow[T \rightarrow \infty]{\text{a.s.}} (0, 0)$ ,
- and asymptotically normal:

$$\sqrt{T} (\alpha_T^{\pm} - a_{\pm}, \beta_T^{\pm} - b_{\pm}) \xrightarrow[T \rightarrow \infty]{\text{law}} N^{\pm} := (N^{\pm, \alpha}, N^{\pm, \beta}),$$

where  $N^+$ ,  $N^-$  are two mutually independent, independent of  $X$ , two-dimensional Gaussian r.v. with covariance matrices given by

$$2\sigma_{\pm} \Gamma_{\pm}^{-1} \text{ with } \Gamma_{\pm} = \begin{pmatrix} Q_{\infty}^{\pm, -1} & -Q_{\infty}^{\pm, 0} \\ -Q_{\infty}^{\pm, 0} & Q_{\infty}^{\pm, 1} \end{pmatrix}.$$

# Drift Estimation from discrete observations

We define  $X_i := X_{t_i}$ , with  $i = 0, \dots, N$  and set

$\Delta_N = \max_{k=0, \dots, N-1} \{t_k - t_{k+1}\}$ . For  $m = -1, 0, 1$  and  $\pm \in \{-, +\}$ , let

$$\mathcal{Q}_{T,N}^{\pm,m} = \sum_{i=0}^{N-1} X_i^m 1_{\{\pm(X_i - r) \geq 0\}} (t_{k+1} - t_k),$$

$$\mathcal{M}_{T,N}^{\pm,m} = \sum_{i=0}^{N-1} X_i^m 1_{\{\pm(X_i - r) \geq 0\}} (X_{i+1} - X_i).$$

We refer with discretized likelihood to

$$G_{T,N}(a_+, b_+, a_-, b_-) = \exp \left( \sum_{i=0}^{N-1} \frac{a_r(X_i) - b_r(X_i)X_i}{2\sigma_r(X_i)X_i} (X_{i+1} - X_i) - \frac{t_{i+1} - t_i}{4} \frac{(a_r(X_i) - b_r(X_i)X_i)^2}{\sigma_r(X_i)X_i} \right)$$

We denote

$$(\alpha_{T,N}^{\pm}, \beta_{T,N}^{\pm}) = \underset{\theta \in \Theta}{\text{Argmax}} G_{T,N}(\theta).$$

# Theorem MLE: Long time and high frequency

Assume  $X_0 \sim \mu$ , where  $\mu$  is the stationary distribution. We suppose that

$$\lim_{N \rightarrow \infty} T_N = +\infty \quad \text{and} \quad \lim_{N \rightarrow \infty} \Delta_N = 0.$$

Let  $\pm \in \{+, -\}$ . It holds that

- The estimator is consistent  $(\alpha_{T_N, N} - a_{\pm}, \beta_{T_N, N} - b_{\pm}) \xrightarrow[N \rightarrow \infty]{\mathbb{P}} (0, 0)$ ,
- and asymptotically normal if  $\lim_{N \rightarrow \infty} T_N \Delta_N = 0$ :

$$\sqrt{T_N} (\alpha_{T_N, N} - a_{\pm}, \beta_{T_N, N} - b_{\pm}) \xrightarrow[N \rightarrow \infty]{\text{law}} N^{\pm} := (N^{\pm, \alpha}, N^{\pm, \beta}),$$

where  $N^+$ ,  $N^-$  are two mutually independent, independent of  $X$ , two-dimensional Gaussian r.v. with covariance matrices given by

$$2\sigma_{\pm} \Gamma_{\pm}^{-1} \text{ with } \Gamma_{\pm} = \begin{pmatrix} Q_{\infty}^{\pm, -1} & -Q_{\infty}^{\pm, 0} \\ -Q_{\infty}^{\pm, 0} & Q_{\infty}^{\pm, 1} \end{pmatrix}.$$

# Sketch of the proof

For all  $N \in \mathbb{N}$  it holds

$$\left( \alpha_{T_N, N}^{\pm} - a_{\pm}, \beta_{T_N, N}^{\pm} - b_{\pm} \right) = \left( \alpha_{T_N, N}^{\pm} - \alpha_{T_N}^{\pm}, \beta_{T_N, N}^{\pm} - \beta_{T_N}^{\pm} \right) + \underbrace{\left( \alpha_{T_N}^{\pm} - a_{\pm}, \beta_{T_N}^{\pm} - b_{\pm} \right)}_{\substack{\text{Continuous time-observations} \\ \text{Theorem}}}.$$

Making use of the Theorem on continuous time observations, we have:

$$(\alpha_{T_N}^{\pm} - a_{\pm}, \beta_{T_N}^{\pm} - b_{\pm}) \xrightarrow[N \rightarrow \infty]{\text{a.s.}} (0, 0),$$

$$\sqrt{T_N}(\alpha_{T_N}^{\pm} - a_{\pm}, \beta_{T_N}^{\pm} - b_{\pm}) \xrightarrow[N \rightarrow \infty]{\text{law}} N^{\pm} := (N^{\pm, \alpha}, N^{\pm, \beta}),$$

We need to verify that

$$\left( \alpha_{T_N, N}^{\pm} - \alpha_{T_N}^{\pm}, \beta_{T_N, N}^{\pm} - \beta_{T_N}^{\pm} \right) \xrightarrow[N \rightarrow \infty]{\mathbb{P}} 0. \implies \text{Consistency} \quad (4)$$

$$\sqrt{T_N} \left( \alpha_{T_N, N}^{\pm} - \alpha_{T_N}^{\pm}, \beta_{T_N, N}^{\pm} - \beta_{T_N}^{\pm} \right) \xrightarrow[N \rightarrow \infty]{\mathbb{P}} 0. \implies \text{Asymptotic normality} \quad (5)$$

We can prove that (4) is verify ( $\approx$  same for (5)).

### Lemma (Consistency)

Let  $\lim_{N \rightarrow \infty} T_N \rightarrow \infty$  and  $\lim_{N \rightarrow \infty} \Delta_N = 0$ . Then for all  $j \in \{-1, 0\}$ ,  $m \in \{-1, 0, 1\}$  it holds

$$\lim_{N \rightarrow \infty} \mathbb{E}_\mu \left[ |Q_{T_N}^{\pm, m} - Q_{T_N, N}^{\pm, m}| \right] = 0 \text{ and } \lim_{N \rightarrow \infty} \mathbb{E}_\mu \left[ |\mathcal{M}_{T_N}^{\pm, j} - \mathcal{M}_{T_N, N}^{\pm, j}| \right] = 0$$

If  $X_0 = x_0 \in \mathbb{R}^+$  p.s., we need to verify:

$$\lim_{N \rightarrow \infty} \mathbb{E}_{x_0} \left[ |Q_{T_N}^{\pm, m} - Q_{T_N, N}^{\pm, m}| \right] = 0 \text{ and } \lim_{N \rightarrow \infty} \mathbb{E}_{x_0} \left[ |\mathcal{M}_{T_N}^{\pm, j} - \mathcal{M}_{T_N, N}^{\pm, j}| \right] = 0$$

→ Transition probability of the TCIR remains unknown.

- 1 Introduction
- 2 Cox–Ingersoll–Ross (CIR) model
  - Model
  - Literature
- 3 Threshold Cox-Ingersoll-Ross (TCIR) model
  - Existence of a strong solution and property
  - Drift Estimations from continuous observations
  - Drift Estimations from high frequency and long time observations
- 4 Conclusion and Opening

- ① We prove the consistency and the asymptotic normal property for the MLE of the parameters  $(a_{\pm}, b_{\pm})$  in the positive recurrent case, and for:
  - Continuous time Observations.
  - High frequency with long time with  $X_0 \sim \mu$ .

**Merci pour votre attention.**