

Parameters estimation of a Threshold CIR process

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Introduction

The Cox–Ingersoll–Ross (CIR) model is a stochastic process that exhibits mean-reverting behaviour.

- In finance, the CIR is used to describes the evolution of interest rates.
- In biology, the CIR can be used to model population dynamics [Bansaye & Méléard, 2015].

Threshold Diffusion:

- Threshold autoregressive (TAR) models in discrete time were introduced in the early 1980s.
- Many applications in Finance, Physics and meteorology.

The Threshold CIR follows the CIR dynamics when above or below a fixed level, yet at this level (threshold) its coefficients can be discontinuous.

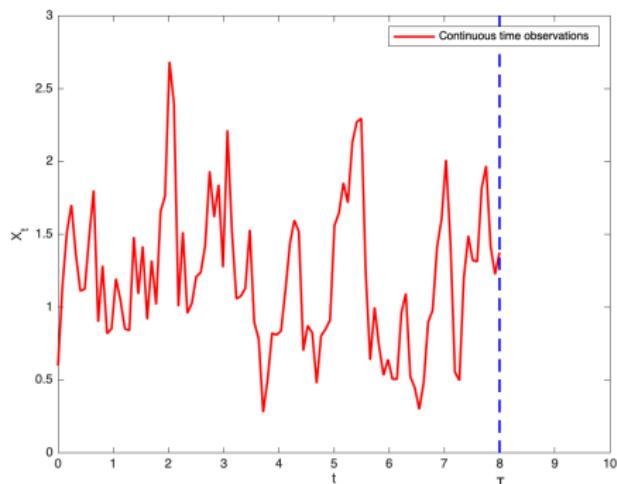
Content

Goal: Estimate the parameters of a TCIR process by using the observations of **a single trajectory**.

- MLE drift estimation.

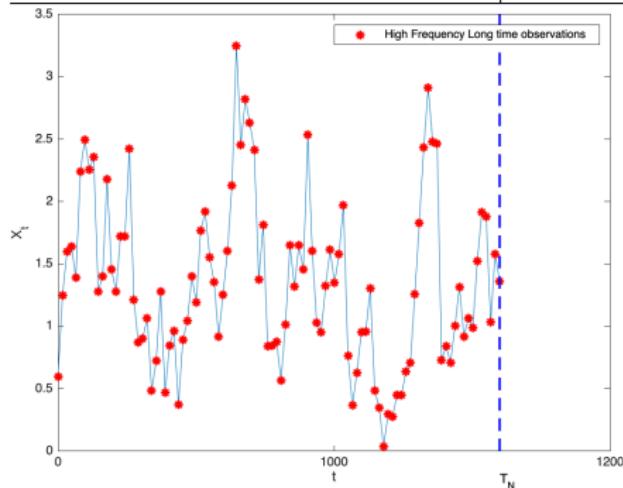
→ Study of the asymptotic behavior of the estimators (Consistency, Asymptotic Normality).

- Continuous time observations.
- High frequency and long time observations.



Continuous time observations:

We observe the process on the interval $[0, T]$ with $T \in (0, \infty)$.



High Frequency in Long time observations:

We observe the process on the discrete time grid $0 < t_0 < \dots < t_N = T$, for $N \in \mathbb{N}$ and $T \in (0, \infty)$.

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We work on the probability space $(\Omega, (\mathcal{F}_t)_{t \geq 0}, \mathbb{P})$.

Cox–Ingersoll–Ross (CIR) process

Let $(X_t)_{t \geq 0}$, the process solution of

$$\begin{cases} dX_t = (a - bX_t)dt + \sqrt{2\sigma X_t}dB_t, & t > 0, \\ X_0 = x_0, \end{cases} \quad (1)$$

with $(a, b, \sigma) \in \mathbb{R}_+ \times \mathbb{R} \times \mathbb{R}_+^*$, $(B_t)_{t \geq 0}$ a standard Brownian motion.

Strong existence of a unique solution to (1) follows from the classical results in 1-D.

Property

- For all $t \geq 0$, the process $(X_t)_{t \geq 0}$ is positif (Comparison Theorem).
- If $a \leq \sigma$, $\{0\}$ instantaneously reflecting.
- If $a > \sigma$, the state space of the process is $]0, +\infty[$.

Literature

In [Alaya & Kebaier, 2013], the authors study the Maximum Likelihood Estimator for the parameters (a, b) .

Likelihood ration/Girsanov Weight

Suppose that the one dimensional diffusion process $(X_t^\theta)_{t \geq 0}$ satisfies

$$dX_t^\theta = b(\theta, X_t)dt + \sigma(X_t)dB_t, \quad X_0 = x_0.$$

- Continuous observations $(X_t^{\theta_0})_{t \in [0, T]}$.
- $\theta_0 \in \Theta$ the parameter to be estimated.

For any $\theta \in \Theta$, the Likelihood ratio is given by

$$G_T^{\theta, \theta_0} := \frac{d\mathbb{P}_{\theta, T}}{d\mathbb{P}_{\theta_0, T}} = \exp \left(\int_0^T \frac{b(\theta, X_s) - b(\theta_0, X_s)}{\sigma^2(X_s)} dX_s - \frac{1}{2} \int_0^T \frac{b^2(\theta, X_s) - b^2(\theta_0, X_s)}{\sigma^2(X_s)} ds \right).$$

For the CIR process the Likelihood ratio, evaluated at time T is given by

$$G_T(\theta) = G_T(a, b) := G_T^{\theta, \theta_0} = \exp \left(\frac{1}{2\sigma} \int_0^T \frac{a-bX_s}{X_s} dX_s - \frac{1}{4\sigma} \int_0^T \frac{(a-bX_s)^2}{X_s} ds \right).$$

$$(\hat{a}_T, \hat{b}_T) = \underset{(a,b) \in \Theta}{\operatorname{Argmax}} G_T(a, b) \text{ (Explicitly computable).}$$

Remark

The estimators and the Likelihood ratio is well defined iff $a > \sigma$ i.e.

$$\mathbb{P}_\theta \left(\int_0^T \frac{ds}{X_s} < \infty \right) = 1.$$

So $\Theta \in \{(a, b) \text{ s.a. } a > \sigma\}$.

Study of the asymptotic behavior in the ergodic case ($b > 0$) and non ergodic case ($b \leq 0$).

- in long time from continuous time observations.
- in long time from high frequency observations (for any $x_0 \in \mathbb{R}^+$).

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Threshold Cox-Ingersoll-Ross (TCIR) model

Let $(X_t)_{t \geq 0}$, the process solution of

$$\begin{cases} dX_t = (a_r(X_t) - b_r(X_t)X_t)dt + \sqrt{2\sigma_r(X_t)X_t}dB_t, & t > 0, \\ X_0 = x_0. \end{cases} \quad (2)$$

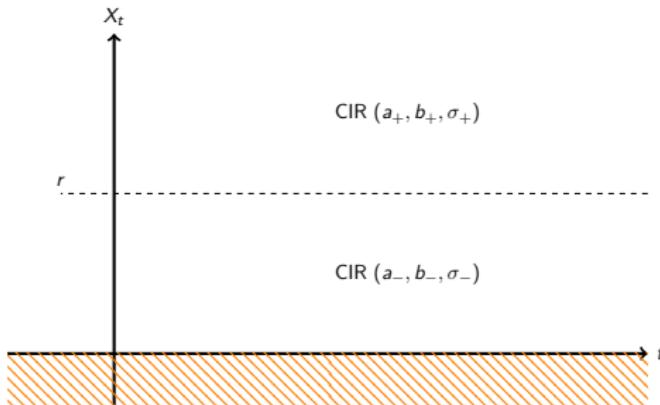
$$\begin{aligned} a_r(x) &= \begin{cases} a_+ \in \mathbb{R} & \text{si } x \geq r, \\ a_- \in \mathbb{R}^+ & \text{si } x < r, \end{cases} & b_r(x) &= \begin{cases} b_+ \in \mathbb{R} & \text{si } x \geq r, \\ b_- \in \mathbb{R} & \text{si } x < r, \end{cases} \\ \sigma_r(x) &= \begin{cases} \sigma_+ > 0 & \text{si } x \geq r, \\ \sigma_- > 0 & \text{si } x < r. \end{cases} \end{aligned} \quad (3)$$

with $r > 0$ et $x_0 > 0$.

Issue: Drift and volatility does not fit into the classical hypotheses of strong existence and unicity of a dimensional EDS.

→ Strong existence of a unique solution to (1) follows from the results of [Mazzonetto & Nieto, in progress].

Property of the TCIR



Separately on (r, ∞) and $(0, r)$, the process follows the CIR dynamics.

Property

- For all $t \geq 0$, the process $(X_t)_{t \geq 0}$ is positif (Comparison Theorem [Mazzonetto & Nieto, in progress]).
- If $a_- \leq \sigma_-$, $\{0\}$ instantaneously reflecting.
- If $a_- > \sigma_-$, the state space of the process is $]0, +\infty[$.

Regime of the process

We denote the scale function S and speed measure m :

$$S'(x) = \exp \left(\int^x \frac{(a_r(y) - b_r(y)y)}{\sigma_r(y)y} dy \right) \text{ and } m(x) = \frac{1}{\sigma_r(x)xS'(x)}.$$

Long time behavior depends on the parameters:

	Null recurrent	Positive Recurrent (Ergodic)	Transience
{0} Absorbing point ($a_- = 0$)	$[b_+ = 0 \text{ and } 0 \leq a_+ < \sigma_+]$ and $[b_- \in \mathbb{R} \text{ and } a_- \leq 0]$	\emptyset	$[b_+ < 0 \text{ or } a_+ \geq \sigma_+]$
{0} Instantaneous reflecting point ($a_- \leq \sigma_-$)	$[b_+ = 0 \text{ and } 0 \leq a_+ < \sigma_+]$ and $[b_- \in \mathbb{R} \text{ and } 0 < a_- \leq \sigma_-]$.	$[[b_+ > 0 \text{ and } a_+ \in \mathbb{R}]]$ or $[[b_+ = 0 \text{ and } a_+ \in \mathbb{R}_-^*]]$ and $[b_- \in \mathbb{R} \text{ and } 0 < a_- \leq \sigma_-]$.	$[b_+ < 0 \text{ or } a_+ \geq \sigma_+]$
{0} is not reached ($a_- > \sigma_-$)	$[b_+ = 0 \text{ and } 0 \leq a_+ < \sigma_+]$ and $[b_- \in \mathbb{R} \text{ and } a_- > \sigma_-]$	$[[b_+ > 0 \text{ and } a_+ \in \mathbb{R}]]$ or $[[b_+ = 0 \text{ and } a_+ \in \mathbb{R}_-^*]]$ and $[b_- \in \mathbb{R} \text{ and } a_- > \sigma_-]$.	$[b_+ < 0 \text{ or } a_+ \geq \sigma_+]$

In the positive recurrent case, the stationary measure μ , is given by

$$\mu(dx) = \frac{m(x)}{\int_0^{+\infty} m(y)dy} dx.$$

We denote $\theta = (a_+, b_+, a_-, b_-)$, the likelihood ratio evaluated at time T is given by

$$G_T(\theta) = \exp \left(\int_0^T \frac{a_r(X_s) - b_r(X_s)X_s}{2\sigma_r(X_s)X_s} dX_s - \frac{1}{4} \int_0^T \frac{(a_r(X_s) - b_r(X_s)X_s)^2}{\sigma_r(X_s)X_s} ds \right).$$

Remark

The estimators and the Likelihood is well defined iff $a_- > \sigma_-$ i.e.

$$\mathbb{P}_\theta \left(\int_0^T \frac{1_{X_s \leq r} ds}{X_s} < \infty \right) = 1.$$

For the estimation problem, we suppose r and σ_\pm to be known.

Drift estimation from continuous time observations

Imagine that we observe $(X_t)_{t \in [0, T]}$ in continuous time. For $T \in (0, \infty)$ and $m = -1, 0, 1$, we define

$$\mathcal{Q}_T^{\pm, m} = \int_0^T X_s^m 1_{\{\pm(X_s - r) \geq 0\}} ds \quad \text{and} \quad \mathcal{M}_T^{\pm, m} = \int_0^T X_s^m 1_{\{\pm(X_s - r) \geq 0\}} dX_s.$$

Let $\theta_0 = (a_0^\pm, b_0^\pm)$, the parameter to be estimated and $\theta \in \Theta \subset \mathbb{R}^4$. We denote

$$(\alpha_T^\pm, \beta_T^\pm) = \underset{\theta \in \Theta}{\operatorname{Argmax}} \ G_T(\theta).$$

Estimators

For every $T \in (0, \infty)$ the MLE are given by

$$\alpha_T^\pm = \frac{\mathcal{M}_T^{\pm, -1} \mathcal{Q}_T^{\pm, 1} - \mathcal{M}_T^{\pm, 0} \mathcal{Q}_T^{\pm, 0}}{\mathcal{Q}_T^{\pm, 1} \mathcal{Q}_T^{\pm, -1} - (\mathcal{Q}_T^{\pm, 0})^2} \quad \text{and} \quad \beta_T^\pm = \frac{\mathcal{M}_T^{\pm, -1} \mathcal{Q}_T^{\pm, 0} - \mathcal{Q}_T^{\pm, -1} \mathcal{M}_T^{\pm, 0}}{\mathcal{Q}_T^{\pm, 1} \mathcal{Q}_T^{\pm, -1} - (\mathcal{Q}_T^{\pm, 0})^2},$$

Theorem MLE: long-time behavior (positive recurrent cases)

Let $\pm \in \{+, -\}$. It holds that

- $\frac{1}{T} \mathcal{Q}_T^{\pm, m} \xrightarrow[T \rightarrow \infty]{\text{a.s.}} \mathcal{Q}_{\infty}^{\pm, m} \in \mathbb{R}_+^*$ (we have explicit expressions using Ergodic Theorem).
- The estimator is strongly consistent $(\alpha_T^{\pm} - a_{\pm}, \beta_T^{\pm} - b_{\pm}) \xrightarrow[T \rightarrow \infty]{\text{a.s.}} (0, 0)$,
- and asymptotically normal:

$$\sqrt{T} (\alpha_T^{\pm} - a_{\pm}, \beta_T^{\pm} - b_{\pm}) \xrightarrow[T \rightarrow \infty]{\text{law}} N^{\pm} := (N^{\pm, \alpha}, N^{\pm, \beta}),$$

where N^+ , N^- are two mutually independent, independent of X , two-dimensional Gaussian r.v. with covariance matrices given by

$$2\sigma_{\pm} \Gamma_{\pm}^{-1} \text{ with } \Gamma_{\pm} = \begin{pmatrix} \mathcal{Q}_{\infty}^{\pm, -1} & -\mathcal{Q}_{\infty}^{\pm, 0} \\ -\mathcal{Q}_{\infty}^{\pm, 0} & \mathcal{Q}_{\infty}^{\pm, 1} \end{pmatrix}.$$

Drift Estimation from discrete observations

We define $X_i := X_{t_i}$, with $i = 0, \dots, N$ and set

$\Delta_N = \max_{k=0 \dots, N-1} \{t_k - t_{k+1}\}$. For $m = -1, 0, 1$ and $\pm \in \{-, +\}$, let

$$\mathcal{Q}_{T,N}^{\pm,m} = \sum_{i=0}^{N-1} X_i^m 1_{\{\pm(X_i - r) \geq 0\}} (t_{k+1} - t_k),$$

$$\mathcal{M}_{T,N}^{\pm,m} = \sum_{i=0}^{N-1} X_i^m 1_{\{\pm(X_i - r) \geq 0\}} (X_{i+1} - X_i).$$

We refer with discretized likelihood to

$$G_{T,N}(a_+, b_+ -, a_-, b_-) = \exp \left(\sum_{i=0}^{N-1} \frac{a_r(X_i) - b_r(X_i)X_i}{2\sigma_r(X_i)X_i} (X_{i+1} - X_i) - \frac{t_{i+1} - t_i}{4} \frac{(a_r(X_i) - b_r(X_i)X_i)^2}{\sigma_r(X_i)X_i} \right)$$

We denote

$$(\alpha_{T,N}^{\pm}, \beta_{T,N}^{\pm}) = \operatorname{Argmax}_{\theta \in \Theta} G_{T,N}(\theta).$$

Theorem MLE: Long time and high frequency

Assume $X_0 \sim \mu$, where μ is the stationary distribution. We suppose that

$$\lim_{N \rightarrow \infty} T_N = +\infty \quad \text{and} \quad \lim_{N \rightarrow \infty} \Delta_N = 0.$$

Let $\pm \in \{+, -\}$. It holds that

- The estimator is consistent $(\alpha_{T_N, N} - a_{\pm}, \beta_{T_N, N} - b_{\pm}) \xrightarrow[N \rightarrow \infty]{\mathbb{P}} (0, 0)$,
- and asymptotically normal if $\lim_{N \rightarrow \infty} T_N \Delta_N = 0$:

$$\sqrt{T_N} (\alpha_{T_N, N} - a_{\pm}, \beta_{T_N, N} - b_{\pm}) \xrightarrow[N \rightarrow \infty]{\text{law}} N^{\pm} := (N^{\pm, \alpha}, N^{\pm, \beta}),$$

where N^+ , N^- are two mutually independent, independent of X , two-dimensional Gaussian r.v. with covariance matrices given by

$$2\sigma_{\pm} \Gamma_{\pm}^{-1} \text{ with } \Gamma_{\pm} = \begin{pmatrix} \mathcal{Q}_{\infty}^{\pm, -1} & -\mathcal{Q}_{\infty}^{\pm, 0} \\ -\mathcal{Q}_{\infty}^{\pm, 0} & \mathcal{Q}_{\infty}^{\pm, 1} \end{pmatrix}.$$

Sketch of the proof

For all $N \in \mathbb{N}$ it holds

$$(\alpha_{T_N, N}^\pm - a_\pm, \beta_{T_N, N}^\pm - b_\pm) = (\alpha_{T_N, N}^\pm - \alpha_{T_N}^\pm, \beta_{T_N, N}^\pm - \beta_{T_N}^\pm) + \underbrace{(\alpha_{T_N}^\pm - a_\pm, \beta_{T_N}^\pm - b_\pm)}_{\substack{\text{Continuous time-observations} \\ \text{Theorem}}},$$

Making use of the Theorem on continuous time observations, we have:

$$(\alpha_{T_N}^\pm - a_\pm, \beta_{T_N}^\pm - b_\pm) \xrightarrow[N \rightarrow \infty]{\text{a.s.}} (0, 0),$$

$$\sqrt{T_N}(\alpha_{T_N}^\pm - a_\pm, \beta_{T_N}^\pm - b_\pm) \xrightarrow[N \rightarrow \infty]{\text{law}} N^\pm := (N^{\pm, \alpha}, N^{\pm, \beta}),$$

We need to verify that

$$(\alpha_{T_N, N}^\pm - \alpha_{T_N}^\pm, \beta_{T_N, N}^\pm - \beta_{T_N}^\pm) \xrightarrow[N \rightarrow \infty]{\mathbb{P}} 0. \implies \text{Consistency} \quad (4)$$

$$\sqrt{T_N}(\alpha_{T_N, N}^\pm - \alpha_{T_N}^\pm, \beta_{T_N, N}^\pm - \beta_{T_N}^\pm) \xrightarrow[N \rightarrow \infty]{\mathbb{P}} 0. \implies \text{Asymptotic normality} \quad (5)$$

We can prove that (4) is verify (\approx same for (5)).

Lemma (Consistency)

Let $\lim_{N \rightarrow \infty} T_N \rightarrow \infty$ and $\lim_{N \rightarrow \infty} \Delta_N = 0$. Then for all $j \in \{-1, 0\}$, $m \in \{-1, 0, 1\}$ it holds

$$\lim_{N \rightarrow \infty} \mathbb{E}_\mu \left[|\mathcal{Q}_{T_N}^{\pm, m} - \mathcal{Q}_{T_N, N}^{\pm, m}| \right] = 0 \text{ and } \lim_{N \rightarrow \infty} \mathbb{E}_\mu \left[|\mathcal{M}_{T_N}^{\pm, j} - \mathcal{M}_{T_N, N}^{\pm, j}| \right] = 0$$

If $X_0 = x_0 \in \mathbb{R}^+$ p.s., we need to verify:

$$\lim_{N \rightarrow \infty} \mathbb{E}_{x_0} \left[|\mathcal{Q}_{T_N}^{\pm, m} - \mathcal{Q}_{T_N, N}^{\pm, m}| \right] = 0 \text{ and } \lim_{N \rightarrow \infty} \mathbb{E}_{x_0} \left[|\mathcal{M}_{T_N}^{\pm, j} - \mathcal{M}_{T_N, N}^{\pm, j}| \right] = 0$$

→ Transition probability of the TCIR remains unknown.

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Conclusion

① We prove the consistency and the asymptotic normal property for the MLE of the parameters (a_{\pm}, b_{\pm}) in the positive recurrent case, and for:

- Continuous time Observations.
- High frequency with long time **with $X_0 \sim \mu$** .

Merci pour votre attention.